

Outline

Test of Coefficients

Intervals for Coefficients

Intervals for π at specific X

Hypothesis Test for β_1

In linear regression, we computed the **test statistic**:

$$t_{obs} = \frac{\hat{\beta}_1 - 0}{\hat{se}(\hat{\beta}_1)}$$

(number of standard errors $\hat{\beta}_1$ is from 0).

P-value: **prob. of getting a test stat this big by chance if H_0 true (i.e., $\beta_1 = 0$)**

Hypothesis Test for β_1

In logistic regression we can do the same thing, but with Normal instead of t distribution.

$$z_{obs} = \frac{\hat{\beta}_1 - 0}{\hat{se}(\hat{\beta}_1)}$$

and get P -value: **prob of a test stat this big if H_0 true**

In R

```
library(Stat2Data); data(MedGPA)
mcatModel <- glm(Acceptance ~ MCAT, data = MedGPA, family = "binomial")
summary(mcatModel) %>% coef() %>% round(3)
```

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-8.712	3.236	-2.692	0.007
MCAT	0.246	0.089	2.752	0.006

Only 0.6% chance we'd get $\left| \hat{\beta}_1 \right| \geq 0.246$ if the association is due solely to chance sampling

Linear vs. Logistic Regression

Goal	Linear	Logistic
Estimate coefs	Minimize SSE	Maximize Likelihood (or Minimize Deviance)
Check conditions	Linearity/Const. var.: Residual vs. Fitted Normality: QQ Plots	Logit linearity: Binned residuals vs. fitted
Test coefs	Measure SEs from 0, P -value using t	Measure SEs from 0 P -value using Normal

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Confidence Interval for β_1

Same principle applies for confidence interval...

$$CI(\Delta\text{logit}) : \hat{\beta}_1 \pm z^* \cdot \hat{se}(\hat{\beta}_1)$$

```
confint(mcatModel) %>% round(2)
```

	2.5 %	97.5 %
(Intercept)	-15.77	-3.04
MCAT	0.09	0.44

- But... β_1 is the rate of change of the log odds, which is hard to understand.
- More common to report a CI for **odds ratio** (e^{β_1}).

$$CI(OR) : (e^{\beta_1^{(lwr)}}, e^{\beta_1^{(upr)}})$$

In R...

```
confint(mcatModel) %>% round(2)
```

	2.5 %	97.5 %
(Intercept)	-15.77	-3.04
MCAT	0.09	0.44

```
confint(mcatModel) %>% exp() %>% round(2)
```

	2.5 %	97.5 %
(Intercept)	0.00	0.05
MCAT	1.09	1.55

“We are 95% confident that the **odds (not probability)** of admittance increases by a **factor of (is multiplied by)** between 1.09 and 1.55 for each additional point of MCAT score”

Linear vs. Logistic Regression

Goal	Linear	Logistic
Estimate coefs	Minimize SSE	Maximize Likelihood
Check conditions	<i>Linearity/Const. var.:</i> Residual vs. Fitted <i>Normality:</i> QQ Plots	<i>Logit linearity:</i> Binned residuals vs. fitted
Test coefs	Measure SEs from 0, <i>P</i> -value using <i>t</i>	Measure SEs from 0 <i>P</i> -value using Normal
Intervals for Params	Slope: β_1	Odds Ratio: e^{β_1}

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Intervals for π at specific X

CI at specific values

Arguably easier to interpret, CIs for π at a few specific X values

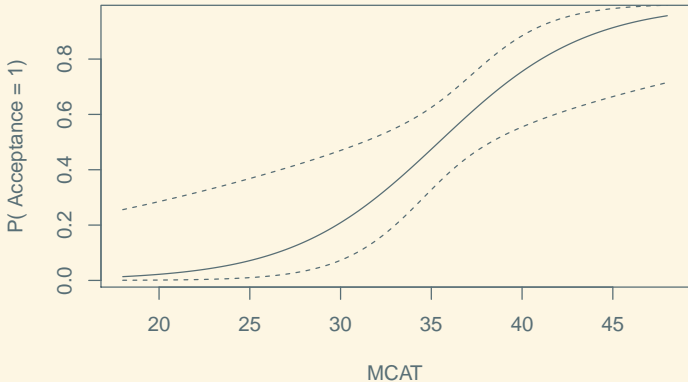
```
source("http://colindawson.net/stat213/code/helper_functions.R")
## functions made with regular makeFun() give point values but not
## intervals with logistic models, so I wrote a custom function
f_hat <- makeFun.logistic(mcatModel)
quartiles <- quantile(~MCAT, data = MedGPA)
f_hat(MCAT = quartiles, interval = "confidence", level = 0.95) %>% round(2)
```

	MCAT	pi.hat	lwr	upr
0%	18	0.01	0.00	0.26
25%	34	0.41	0.26	0.58
50%	36	0.54	0.39	0.67
75%	39	0.71	0.52	0.84
100%	48	0.96	0.72	0.99

Interpretation: “We are 95% confident that the **probability** of acceptance for students with an MCAT score of 39 is between 52% and 84%”

Confidence Bands

```
## Also requires sourcing helper_functions.R  
## Can supply level=, xlim=, xlab= and ylab= to customize graph  
plot.logistic.bands(mcatModel)
```



Linear vs. Logistic Regression

Goal	Linear	Logistic
Estimate coefs	Minimize SSE	Maximize Likelihood
Check conditions	<i>Linearity/Const. var.:</i> Residual vs. Fitted <i>Normality:</i> QQ Plots	<i>Logit linearity:</i> Binned residuals vs. fitted
Test coefs	Measure SEs from 0, <i>P</i> -value using <i>t</i>	Measure SEs from 0 <i>P</i> -value using Normal
Intervals for Params	Slope: β_1	Odds Ratio: e^{β_1}
Intervals for Fitted Vals.	Confidence and prediction intervals	Confidence intervals only